

CLASSICAL MODEL

OLS: Ass: $E(Y|X) = X\beta$ $V(Y|X) = \sigma^2 I_n$ $\Pr[rk(X'X) = k] = 1$

$$\hat{\beta}_{OLS} = (X'X)^{-1} X'Y \quad \hat{Y} = PY \quad \hat{e} = MY \quad X'\hat{e} = 0 \quad \text{nx } \hat{\beta}_1 = (X_1' M_2 X_1)^{-1} X_1' M_2 Y$$

$$r = \frac{\sum (X_i - \bar{X})(Y_i - \bar{Y})}{\sqrt{\sum (X_i - \bar{X})^2} \sqrt{\sum (Y_i - \bar{Y})^2}}, \quad 1x \hat{\beta}_1 = r_{xy} \frac{s_y}{s_x}, \quad R^2 = r^2, \quad 2x \hat{\beta}_2 = \frac{\hat{\beta}_{12} - \hat{\beta}_{13} \hat{\beta}_{23}}{1 - \hat{\beta}_{23} \hat{\beta}_{32}}, \quad \hat{\beta}_{ij} = r_{ij} \frac{s_i}{s_j}$$

$$R_{UC}^2 = 1 - \frac{\hat{e}'\hat{e}}{Y'Y} = \frac{Y'\hat{Y}}{Y'Y} \in [0, 1] \quad R^2 = R_C^2 = 1 - \frac{\hat{e}'\hat{e}}{Y' M_0 Y} \in [0, 1] \quad \text{if c } R_{adj}^2 = 1 - (1 - R^2) \frac{n-1}{n-k}$$

Properties: $\Gamma\hat{\beta} - BLUE : E(\hat{\beta}|X) = \beta, Var(\hat{\beta}|X) = \sigma^2 (X'X)^{-1}, Cov(\hat{\beta}, \hat{e}|X) = 0$

$$Var(\hat{\beta}_1|X) = \sigma^2 (X_1' M_2 X_1)^{-1}, \hat{\sigma}_{OLS}^2 = \frac{\hat{e}'\hat{e}}{n-k} - \text{unbiased}, V(\hat{Y}_0|X) = \sigma^2 (I + X_0 (X'X)^{-1} X_0')$$

Ass: Joint Normality: $Y|X \sim N_n(X\beta, \sigma^2 I_n) \implies$ RandSamp: $(X, Y)_i \sim i.i.d.$

$$\Gamma\hat{\beta}|X \sim N_k(\Gamma\beta, \sigma^2 \Gamma (X'X)^{-1} \Gamma'), W = (\Gamma\hat{\beta} - \Gamma\beta)' (\sigma^2 \Gamma (X'X)^{-1} \Gamma')^{-1} (\Gamma\hat{\beta} - \Gamma\beta) \sim \chi^2(rk(\Gamma))$$

$$\frac{\Gamma\hat{\beta} - \Gamma\beta}{\sqrt{\hat{\sigma}^2 \Gamma (X'X)^{-1} \Gamma'}} \sim t(rk(\Gamma)), \quad \frac{(\Gamma\hat{\beta} - \Gamma\beta)' (\hat{\sigma}^2 \Gamma (X'X)^{-1} \Gamma')^{-1} (\Gamma\hat{\beta} - \Gamma\beta)}{rk(\Gamma)} \sim F(rk(\Gamma), n-k)$$

HyD SignLevel = $\Pr(\text{Rejected}|\text{null}) = \Pr(I \text{ error}),$ Power = $\Pr(\text{Rejected}|\text{alter}) = 1 - \Pr(II \text{ error})$

$$R : \Gamma\hat{\beta} = \gamma, F_0 = \frac{RSS_R - RSS_U}{RSS_U} \frac{n-k}{p} = \frac{R_U^2 - R_R^2}{1 - R_U^2} \frac{n-k}{p}, \hat{\beta}_R = \hat{\beta}_U - (X'X)^{-1} \Gamma' (\Gamma (X'X)^{-1} \Gamma')^{-1} (\Gamma\hat{\beta} - \Gamma\beta)$$

MLE $\hat{\beta}_{ML} = \hat{\beta}_{OLS} = (X'X)^{-1} X'Y,$ attains CLRB, $\hat{\sigma}_{ML}^2 = \frac{\hat{e}'\hat{e}}{n}$ - biased, does not

$$I(\beta, \sigma^2|X)^{-1} = -E \left[\frac{\partial^2 L}{\partial \theta \partial \theta'} | X \right]^{-1} = \begin{bmatrix} \sigma^2 (X'X)^{-1} & 0 \\ 0 & 2\sigma^4/n \end{bmatrix}, \quad \sqrt{n} (\hat{\theta}_{ML} - \theta) \xrightarrow{d} N(0, \lim \frac{1}{n} I_n^{-1})$$

$$\text{Tst: } \frac{W}{n} = h(\theta)' \left\{ \frac{\partial Q}{\partial \theta} \left[\frac{\partial^2 Q}{\partial \theta \partial \theta'} \right]^{-1} \frac{\partial Q}{\partial \theta} \right\}^{-1} h(\theta) \Big|_{\hat{\theta}_U}, \quad \frac{LM}{n} = \frac{\partial Q}{\partial \theta} \left[\frac{\partial^2 Q}{\partial \theta \partial \theta'} \right]^{-1} \frac{\partial Q}{\partial \theta} \Big|_{\hat{\theta}_R}, \quad \frac{LR}{n} = 2 \left(Q(\hat{\theta}_R) - Q(\hat{\theta}_U) \right)$$

$$\text{CNLR: } W = n \frac{RSS_R - RSS_U}{RSS_U}, \quad LM = n \frac{RSS_R - RSS_U}{RSS_R}, \quad LR = n \ln \frac{RSS_R}{RSS_U}, \quad LM \leq LR \leq W, \text{ all } \sim \chi^2(rk(\Gamma))$$

Deviations Multicol: $Var(\hat{\beta}_j|X) = \sigma^2 (X_j' M_{-j} X_j)^{-1} = \frac{\sigma^2}{(1 - R_j^2)(X_j' M_j X_j)}$

Redund: no bias **Omit:** bias: $(X_1' X_1)^{-1} X_1 X_2 \beta_2,$ Both: $V_{true} \uparrow, V_{est} ?$

Heter: $V(X), Var(\hat{\beta}|X) = \sigma^2 (X'X)^{-1} X'V(X) X (X'X)^{-1}, \hat{\beta}_{GLS} = (X'VX)^{-1} X'VY$ - BLUE

$$\text{AR(1): } \varepsilon_i = \phi \varepsilon_{i-1} + u_i \quad Y_i - \phi Y_{i-1} = (X_i - \phi X_{i-1}) \beta + u_i \quad Y_1 = \sqrt{1 - \phi^2} X_1 \beta + u_1$$

ASYMPTOTICS

Asy prop OLS Ass: $Y_i = X_i \beta + \varepsilon_i, \{X_i, \varepsilon_i\} i.i.d., E[X_i \varepsilon_i] = 0, E|X_i \varepsilon_i| < \infty, E|X_i|^2 < \infty, E[X_i' X_i] p.s.d.$

Consistent: $\hat{\beta} = \beta + \left(\frac{1}{n} \sum X_i' X_i \right)^{-1} \left(\frac{1}{n} \sum X_i' \varepsilon_i \right) \xrightarrow{p} \beta,$ Ass: $same + E|X_i \varepsilon_i|^2 < \infty, E[X_i' X_i \varepsilon_i^2] p.s.d.$

$$\sqrt{n} (\hat{\beta} - \beta) = \sqrt{n} \left(\frac{1}{n} \sum X_i' X_i \right)^{-1} \left(\frac{1}{n} \sum X_i' \varepsilon_i \right) \xrightarrow{d} N(0, E[X_i' X_i]^{-1} E[X_i' X_i \varepsilon_i^2] E[X_i' X_i]^{-1})$$

i.i.d.: $\hat{\beta} \overset{A}{\sim} N\left(0, \frac{1}{n} AV(\hat{\beta})\right), \frac{1}{n} AV(\hat{\beta}) = \frac{1}{n} \left[\frac{1}{n} \sum X_i' X_i \right]^{-1} \left[\frac{1}{n} \sum X_i' X_i \varepsilon_i^2 \right] \left[\frac{1}{n} \sum X_i' X_i \right]^{-1}$ - **White**

$$t_0 = \frac{\Gamma\hat{\beta} - \Gamma\beta_0}{\sqrt{\Gamma \frac{1}{n} AV(\hat{\beta}) \Gamma'}} \xrightarrow{d} N(0, 1) \quad F_0 = (\Gamma\hat{\beta} - \Gamma\beta_0)' \left(\Gamma \frac{1}{n} AV(\hat{\beta}) \Gamma' \right)^{-1} (\Gamma\hat{\beta} - \Gamma\beta_0) \xrightarrow{d} \chi^2(rk(\Gamma))$$

$$W_0 = (\gamma(\hat{\beta}) - \gamma(\beta))' \left(\Gamma(\hat{\beta}) \frac{1}{n} AV(\hat{\beta}) \Gamma(\hat{\beta})' \right)^{-1} (\gamma(\hat{\beta}) - \gamma(\beta)) \xrightarrow{d} \chi^2(rk(\Gamma))$$
 - **Wald**

Asy WLS: $Y_i = X_i \beta + \varepsilon_i, \{X_i, \varepsilon_i\} i.i.d., E[X_i | \varepsilon_i] = 0,$ etc $\varepsilon_i^2 = Z_i \alpha + u_i, \hat{\varepsilon}_i^2,$

$$\hat{\beta}_{FWLS} = \beta + \left(\frac{1}{n} \sum \frac{X_i' X_i}{\hat{\varepsilon}_i^2} \right)^{-1} \left(\frac{1}{n} \sum \frac{X_i' \varepsilon_i}{\hat{\varepsilon}_i^2} \right) \xrightarrow{p} \beta_{WLS} \xrightarrow{p} \beta \text{ asy more efficient than OLS}$$

EXTREMUM ESTIMATORS

Extr. Est: $\hat{\theta} = \arg \max_{\theta \in \Theta} Q_n(\theta)$ *Lem:* $\Theta \text{ compact}, Q_n \text{ continuous measurable}_1 \implies \exists \hat{\theta}$

Th1: $\Theta Q_n \text{ same}_1, \exists Q_0(\theta) : \exists! \theta_0 = \arg \max_{\theta \in \Theta} Q_0(\theta), Q_n \xrightarrow{p} Q_0 \implies \hat{\theta} \xrightarrow{p} \theta_0$

Th2: $\theta_0 \in \Theta \text{ interior}, Q_n \text{ concave meas}, \exists Q_0(\theta) : \exists! \theta_0 = \arg \max Q_0(\theta), Q_n \xrightarrow{p} Q_0 \implies \hat{\theta} \xrightarrow{p} \theta_0$

$$\text{[M-type]}: Q_n(\theta) = \frac{1}{n} \sum m(W_i, \theta)$$

Th3: $X_i \text{ i.i.d.}, \Theta \text{ compact}, m(W, \theta) \in C^0, E[\sup_{\theta \in \Theta} |m(W_i, \theta)|] < \infty \xrightarrow{ULLN} \hat{\theta} \xrightarrow{p} \theta_0$

Th4: $\text{i.i.d.}, \theta_0 \in \Theta \text{ int}, m(W, \theta) \text{ concave}, E[|m(W_i, \theta)|] < \infty \xrightarrow{LLN} \hat{\theta} \xrightarrow{p} \theta_0$

Th5: $\text{i.i.d.}, \theta_0 \in \Theta \text{ int}, m(W, \theta) \in C^2, \hat{\theta} \xrightarrow{p} \theta_0, E\left[\sup \left| \frac{\partial^2 m(W_i, \theta)}{\partial \theta \partial \theta'} \right| \right] < \infty, 0 < E\left[\frac{\partial^2 m(W_i, \theta)}{\partial \theta \partial \theta'} \right] < \infty,$
 $\sum \frac{\partial m(W_i, \theta)}{\sqrt{n} \partial \theta} \xrightarrow{d} N\left(0, E\left[\frac{\partial m(W_i, \theta)}{\partial \theta} \frac{\partial m(W_i, \theta)}{\partial \theta'} \right] \text{ p.s.d.} \right) \xrightarrow{CLT} \sqrt{n}(\hat{\theta} - \theta_0) \xrightarrow{d} N(0, H_0^{-1} \Omega_0 H_0^{-1})$

NONLINEAR LEAST SQUARES

[NLS]: Ass: $Y_i = g(X_i, \beta_0) + \varepsilon_i, \{X_i, \varepsilon_i\} \text{ i.i.d.}, E[\varepsilon_i | X_i] = 0, \beta \neq \beta_0 \implies g(X_i, \beta) \neq g(X_i, \beta_0)$

$m(W_i, \beta) = (Y_i - g(X_i, \beta))^2$, **[Th1]:** $B \text{ comp}, g(X, \beta) \in C^0, E[\sup ||m(W_i, \beta)||] < \infty \implies \hat{\beta} \xrightarrow{p} \beta_0$

$s(W_i, \beta) = -\frac{\partial g(W_i, \beta)}{\partial \beta} (Y_i - g(X_i, \beta)), H(W_i, \beta) = \frac{\partial g(W_i, \beta)}{\partial \beta} \frac{\partial g(W_i, \beta)}{\partial \beta} - \frac{\partial^2 g(W_i, \beta)}{\partial \beta \partial \beta'} (Y_i - g(X_i, \beta))$

[Th2]: $\text{i.i.d.}, \beta_0 \in B \text{ int}, m(W, \beta) \in C^2, E[\sup |H(W_i, \beta)|] < \infty, 0 < E[s(W_i, \beta_0)^2] < \infty,$

$0 < H_0 = E[H(W_i, \beta_0)] < \infty, \hat{\beta} \xrightarrow{p} \beta_0 \xrightarrow{CLT} \sqrt{n}(\hat{\beta} - \beta_0) \xrightarrow{d} N(0, H_0^{-1} \Omega_0 H_0^{-1})$

MAXIMUM LIKELIHOOD

[ML]: Ass: $f(w_i, \theta_0), w_i \text{ i.i.d.}, m(W_i, \beta) = \ln f(w_i, \theta), K\text{-L: } \theta_0 \text{ id} \implies E[\ln f(w_i, \theta)] < E[\ln f(w_i, \theta_0)]$

Th1: $\text{i.i.d.}, \Theta \text{ compact}, \theta_0 \text{ id}, \ln f(w_i, \theta) \in C^0, E[\sup_{\theta \in \Theta} ||\ln f(w_i, \theta)||] < \infty \implies \hat{\theta}_{ML} \xrightarrow{p} \theta_0$

Th2: $\text{i.i.d.}, \theta_0 \in \Theta \text{ int}, \ln f(w_i, \theta) > 0 \in C^2, \int \sup \left| \frac{\partial f(w_i, \theta)}{\partial \theta} \right| dw_i < \infty, \int \sup \left| \frac{\partial^2 f(w_i, \theta)}{\partial \theta \partial \theta'} \right| dw_i < \infty,$

$E\left[\sup \left| \frac{\partial^2 \ln f(w_i, \theta)}{\partial \theta \partial \theta'} \right| \right] < \infty, 0 < i(\theta_0) = -E\left[\left(\frac{\partial^2 \ln f(w_i, \theta)}{\partial \theta \partial \theta'}\right)^2\right] < \infty, \hat{\theta} \xrightarrow{p} \theta_0$

$$\xrightarrow{CLT} \sqrt{n}(\hat{\theta}_{ML} - \theta_0) \xrightarrow{d} N(0, i(\theta_0)^{-1})$$

$$\text{[TESTS]} \frac{W}{n} = h(\theta)' \left\{ \frac{\partial h'}{\partial \theta} \hat{V}^{-1} \frac{\partial h}{\partial \theta} \right\}^{-1} h(\theta) \Big|_{\hat{\theta}_U}, \frac{LM}{n} = \frac{\partial Q_n'}{\partial \theta} \left[\frac{\partial^2 Q_n}{\partial \theta \partial \theta'} \right]^{-1} \frac{\partial Q_n}{\partial \theta} \Big|_{\hat{\theta}_R}, \frac{LR}{n} = 2 \left(Q_n(\hat{\theta}_R) - Q_n(\hat{\theta}_U) \right)$$

$$\text{all } \sim \chi^2(rk(\Gamma)) \quad \text{ex: } Q_n = \frac{1}{n} \sum \ln f(w_i, \theta)$$

SAMPLE SELECTION MODELS

[Qualit. Resp] $\Pr(y_i = 1 | x_i) =:$ LPM: $x_i \beta_0$, Probit: $\int_{-\infty}^{x_i \beta_0} \frac{e^{-\frac{u^2}{2}}}{\sqrt{2\pi}} du$, Logit: $\frac{e^{x_i \beta_0}}{1 + e^{x_i \beta_0}}$, LogWeib: $e^{-e^{x_i \beta_0}}$

$E[y_i | x_i] = F(x_i \beta_0), NLS : \sum (y_i - F(x_i \beta_0))^2, WNLS : \sum \frac{(y_i - F(x_i \beta_0))^2}{F(x_i \beta_0)(1 - F(x_i \beta_0))}$,

$ML : \prod F(x_i \beta_0)^{y_i} (1 - F(x_i \beta_0))^{1 - y_i}, \text{ MargEff: } f(\bar{x} \beta) \beta, \delta\text{-met. Compare: Logit*0.5513}$

[Censored] $y_i = \max[x_i \beta_0 + \varepsilon_i, 0]: \varepsilon_i | x_i \sim N(0, \sigma_0^2) \quad E[y_i | x_i] = \left[x_i \beta_0 + \sigma_0 \frac{\phi(x_i \beta_0)}{\Phi(x_i \beta_0)} \right] \Phi(x_i \beta_0)$

$$\text{Var}[y_i | x_i] = \sigma_0^2 \left[1 + \frac{x_i \beta_0 \phi\left(\frac{x_i \beta_0}{\sigma_0}\right)}{\sigma_0 \Phi\left(\frac{x_i \beta_0}{\sigma_0}\right)} - \left(\frac{\phi\left(\frac{x_i \beta_0}{\sigma_0}\right)}{\Phi\left(\frac{x_i \beta_0}{\sigma_0}\right)} \right)^2 \right] \Phi\left(\frac{x_i \beta_0}{\sigma_0}\right)^2$$

$NLS, WNLS, ML : \prod [1 - \Phi\left(\frac{x_i \beta}{\sigma}\right)]^{1 - d_i} \left(\frac{1}{\sigma} \phi\left(\frac{y_i - x_i \beta}{\sigma}\right)\right)^{d_i}, \text{ Heck2S : } \text{probit} \hat{\gamma} \rightarrow x + \hat{\lambda}(-x \hat{\gamma})$

$\text{MargEff: expected } (1 + \lambda'(-x \gamma_0)) \beta, \text{ realiz: } \beta \Phi(x \gamma_0)$

$$\text{[Truncated]} y_i = x_i \beta_0 + u_i, \varepsilon_i | x_i \sim N(0, \sigma_0^2) \quad f[y_i | x_i] = \frac{\phi\left(\frac{x_i \beta_0}{\sigma_0}\right)}{\sigma_0 \Phi\left(\frac{x_i \beta_0}{\sigma_0}\right)} ML : \prod \frac{\frac{1}{\sigma} \phi\left(\frac{y_i - x_i \beta}{\sigma}\right)}{\Phi\left(\frac{x_i \beta}{\sigma}\right)}$$