

Modes of Convergence

Definition 1 (convergence almost surely). The matrix-valued sequence of random variables Z_n is said to converge to a random matrix Z *almost surely* (or *with probability one*), written as $Z_n \xrightarrow{as} Z$, if

$$\Pr \left\{ \lim_{n \rightarrow \infty} Z_n = Z \right\} = 1,$$

i.e. almost every trajectory converges to Z .

Definition 2 (convergence in probability). The matrix-valued sequence of random variables Z_n is said to converge to a random matrix Z *in probability*, written as $Z_n \xrightarrow{p} Z$ or $p \lim Z_n = Z$, if

$$\forall \varepsilon > 0 \quad \lim_{n \rightarrow \infty} \Pr \{ \|Z_n - Z\| > \varepsilon \} = 0,$$

i.e. the probability of large deviations converges to 0.

Result. $Z_n \xrightarrow{as} Z \Rightarrow Z_n \xrightarrow{p} Z$.

Definition 3 (convergence in mean square). The matrix-valued sequence of random variables Z_n is said to converge to a random matrix Z *in mean square*, written as $Z_n \xrightarrow{ms} Z$, if

$$\lim_{n \rightarrow \infty} \mathbb{E} [\|Z_n - Z\|^2] = 0,$$

i.e. the mean squared error converges to 0.

Result. $Z_n \xrightarrow{ms} Z \Rightarrow Z_n \xrightarrow{p} Z$.

Definition 4 (convergence in distribution). The vector-valued sequence of random variables Z_n is said to converge to a random vector Z *in distribution*, written as $Z_n \xrightarrow{d} Z$ or $Z_n \xrightarrow{d} \mathcal{D}_Z$, where \mathcal{D}_Z is the distribution of Z , if

$$\lim_{n \rightarrow \infty} \Pr\{Z_n \leq z\} = \Pr\{Z \leq z\}$$

for all continuity points z of $\Pr\{Z \leq z\}$.

Result. $Z_n \xrightarrow{p} Z \Rightarrow Z_n \xrightarrow{d} Z$. If Z is constant, $Z_n \xrightarrow{p} Z \Leftrightarrow Z_n \xrightarrow{d} Z$.

Continuous Mapping Theorems

Theorem (Mann–Wald). Suppose that $g(z)$ is a continuous $\mathbb{R}^{k_1 \times k_2} \rightarrow \mathbb{R}^{\ell_1 \times \ell_2}$ function.

- If $Z_n \xrightarrow{as} Z$ as $n \rightarrow \infty$, then $g(Z_n) \xrightarrow{as} g(Z)$.
- If $Z_n \xrightarrow{p} Z$ as $n \rightarrow \infty$, then $g(Z_n) \xrightarrow{p} g(Z)$.
- If $Z_n \xrightarrow{ms} Z$ as $n \rightarrow \infty$ and g is linear, then $g(Z_n) \xrightarrow{ms} g(Z)$.
- If $Z_n \xrightarrow{d} Z$ as $n \rightarrow \infty$, then $g(Z_n) \xrightarrow{d} g(Z)$.

If in addition $Z = \text{const}$, then continuity of g only at Z suffices.

Theorem (Slutsky). If $U_n \xrightarrow{p} U = \text{const}$ and $V_n \xrightarrow{d} V$ as $n \rightarrow \infty$, then

- $U_n + V_n \xrightarrow{d} U + V$.
- $U_n V_n \xrightarrow{d} UV$, $V_n U_n \xrightarrow{d} VU$.
- $U_n^{-1} V_n \xrightarrow{d} U^{-1} V$, $V_n U_n^{-1} \xrightarrow{d} VU^{-1}$ if $\Pr\{\det(U_n) = 0\} = 0$.

Delta Method

Theorem (Delta Method). Let the sequence of $k \times 1$ random vectors Z_n satisfy

$$\sqrt{n}(Z_n - Z) \xrightarrow{d} \mathcal{N}(0, \Sigma),$$

as $n \rightarrow \infty$, where $Z = p \lim Z_n$ is constant, and the $\mathbb{R}^k \rightarrow \mathbb{R}^\ell$ function $g(z)$ be continuously differentiable at Z . Then

$$\sqrt{n}(g(Z_n) - g(Z)) \xrightarrow{d} \mathcal{N}(0, G\Sigma G'),$$

where $G = \left. \frac{\partial g(z)}{\partial z'} \right|_{z=Z}$.

Laws of Large Numbers

Theorem A (Kolmogorov, independent identical observations). Let $\{Z_i\}_{i=1}^{\infty}$ be independent and identically distributed (IID), and let $\mathbb{E}[|Z_i|]$ exist. Then

$$\frac{1}{n} \sum_{i=1}^n Z_i \xrightarrow{as} \mathbb{E}[Z_i]$$

as $n \rightarrow \infty$.

Theorem B (Kolmogorov, independent heterogeneous observations). Let $\{Z_i\}_{i=1}^{\infty}$ be independent with finite variance σ_i^2 . If

$$\sum_{i=1}^{\infty} \frac{\sigma_i^2}{i^2} < \infty,$$

then

$$\frac{1}{n} \sum_{i=1}^n Z_i - \mathbb{E} \left[\frac{1}{n} \sum_{i=1}^n Z_i \right] \xrightarrow{as} 0$$

as $n \rightarrow \infty$.

Theorem C (Chebyshev, uncorrelated observations). Let $\{Z_i\}_{i=1}^{\infty}$ be uncorrelated, i.e. $\mathbb{C}[Z_i, Z_j] = 0$ if $i \neq j$. If

$$\frac{1}{n^2} \sum_{i=1}^n \sigma_i^2 \xrightarrow{n \rightarrow \infty} 0$$

then

$$\frac{1}{n} \sum_{i=1}^n Z_i - \mathbb{E} \left[\frac{1}{n} \sum_{i=1}^n Z_i \right] \xrightarrow{p} 0$$

as $n \rightarrow \infty$.

Theorem D (Birkhoff-Khinchin, dependent observations, "Ergodic Theorem"). Let $\{Z_t\}_{t=1}^{\infty}$ be a stationary and ergodic sequence of random variables, and let $\mathbb{E}[|Z_t|] < \infty$. Then

$$\frac{1}{T} \sum_{t=1}^T Z_t \xrightarrow{as} \mathbb{E}[Z_t]$$

as $T \rightarrow \infty$.

Central Limit Theorems

Theorem E (Lindberg-Levy, independent identical observations). Let $\{Z_i\}_{i=1}^{\infty}$ be independent and identically distributed (IID) with $\mathbb{E}[Z_i] = \mu$ and $\mathbb{V}[Z_i] = \sigma^2$. Then

$$\sqrt{n} \left(\frac{1}{n} \sum_{i=1}^n Z_i - \mu \right) \xrightarrow{d} \mathcal{N}(0, \sigma^2)$$

as $n \rightarrow \infty$.

Theorem F (Lyapunov, independent heterogeneous observations). Let $\{Z_i\}_{i=1}^{\infty}$ be independent with $\mathbb{E}[Z_i] = \mu_i$, $\mathbb{V}[Z_i] = \sigma_i^2$ and $\mathbb{E}[|Z_i - \mu_i|^3] = \nu_i$. If

$$\frac{(\sum_{i=1}^n \nu_i)^{\frac{1}{3}}}{(\sum_{i=1}^n \sigma_i^2)^{\frac{1}{2}}} \xrightarrow{n \rightarrow \infty} 0,$$

then

$$\frac{\sum_{i=1}^n (Z_i - \mu_i)}{(\sum_{i=1}^n \sigma_i^2)^{\frac{1}{2}}} \xrightarrow{d} \mathcal{N}(0, 1)$$

as $n \rightarrow \infty$.

Theorem G (Billingsley, martingale difference sequences). Let $\{Z_t\}_{t=-\infty}^{+\infty}$ be a stationary and ergodic martingale difference sequence relative to its own past, with $\sigma^2 = \mathbb{E}[Z_t^2] < \infty$. Then

$$\frac{1}{\sqrt{T}} \sum_{t=1}^T Z_t \xrightarrow{d} \mathcal{N}(0, \sigma^2)$$

as $T \rightarrow \infty$.

Theorem H (general dependent observations). Let $\{Z_t\}_{t=-\infty}^{+\infty}$ be a stationary and ergodic sequence of random variables with

$$v_z = \sum_{j=-\infty}^{+\infty} \mathbb{C}[Z_t, Z_{t-j}] < \infty.$$

Then under suitable conditions,

$$\sqrt{T} \left(\frac{1}{T} \sum_{t=1}^T Z_t - \mathbb{E}[Z_t] \right) \xrightarrow{d} \mathcal{N}(0, v_z)$$

as $T \rightarrow \infty$.